

Άσκηση 1 (Ίδίο με θέμα 2016 → Ιανουάριος 2017)

Θεωρούμε ως τυχαία μεταβλητή  $f(x_i; \lambda) = \frac{e^{-\lambda} \lambda^{x_i}}{x_i!}, x_i = 0, 1, 2, \dots$   
 $X_1, X_2, X_3, \dots, X_n \in \mathbb{Z} \quad X_i \sim \text{Poisson}(\lambda)$

$E(X_i) = \text{Var}(X_i) = \lambda$

Από γραμμικό παρέλκο  $\rightarrow X_i = \lambda + u_i$

(i)  $\prod_{i=1}^n f(x_i; \lambda) = f(x_1; \lambda) \cdot f(x_2; \lambda) \cdot \dots \cdot f(x_n; \lambda)$

$= \frac{e^{-\lambda} \lambda^{x_1}}{x_1!} \cdot \frac{e^{-\lambda} \lambda^{x_2}}{x_2!} \cdot \dots \cdot \frac{e^{-\lambda} \lambda^{x_n}}{x_n!} = \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!} = L(\lambda; x_i)$

$\ln L = \ln \left[ \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!} \right] = \ln \left[ e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i} \right] - \ln \left( \prod_{i=1}^n x_i! \right)$

$= \ln(e^{-n\lambda}) + \ln(\lambda^{\sum_{i=1}^n x_i}) - \ln \left( \prod_{i=1}^n x_i! \right)$

$= -n\lambda + \sum_{i=1}^n x_i \ln \lambda - \ln \left( \prod_{i=1}^n x_i! \right)$

F.O.C:  $\frac{\partial \ln L}{\partial \lambda} = -n + \frac{1}{\lambda} \sum_{i=1}^n x_i + 0 = 0$

$\Leftrightarrow \frac{1}{\lambda} \sum_{i=1}^n x_i = n \Leftrightarrow \boxed{\hat{\lambda} = \frac{1}{n} \sum_{i=1}^n x_i}$

S.O.C:  $\frac{\partial^2 \ln L}{\partial \lambda^2} = 0 - \frac{1}{\lambda^2} \sum_{i=1}^n x_i = -\frac{1}{\lambda^2} n\lambda = -\frac{n}{\lambda}$

δεα μέγιστο

## (ii) Απεροσυστία

$$\begin{aligned} E(\hat{\lambda}) &= E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n E(x_i) = \\ &= \frac{1}{n} \sum_{i=1}^n \lambda = \frac{1}{n} \cdot n\lambda = \lambda \end{aligned}$$

### Συνεπεία

$$\begin{aligned} \text{Var}(\hat{\lambda}) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(x_i) = \\ &= \frac{1}{n^2} \sum_{i=1}^n \lambda = \frac{1}{n^2} \cdot n\lambda = \frac{\lambda}{n} \end{aligned}$$

### Τραβές συνθίρες

- $\lim_{n \rightarrow \infty} E(\hat{\lambda}) = \lim_{n \rightarrow \infty} \lambda = \lambda \checkmark$
- $\lim_{n \rightarrow \infty} \text{Var}(\hat{\lambda}) = \lim_{n \rightarrow \infty} \frac{\lambda}{n} = 0 \checkmark$

Άρα η στατιστική είναι συνεπεία.

### Μέγιστο αποτελεσματικότητα

$$\begin{aligned} C-R &= I_n^{-1}(\hat{\lambda}) = \left[ E\left(-\frac{\partial^2 \ln L}{\partial \lambda^2}\right) \right]^{-1} = \left[ E\left[-\left(-\frac{n}{\lambda}\right)\right] \right]^{-1} \\ &= \frac{\lambda}{n} = \text{Var}(\hat{\lambda}) \end{aligned}$$

Επομένως η στατιστική είναι η μέγιστο αποτελεσματική.

Άσκηση 2 (IS10. Σεφρα 2017 → Τάσος 2017) (3)

Τυχαίο δείγμα  $X_1, X_2, \dots, X_n$  με  $X_i \sim \text{Bernoulli}(\theta)$

$$E(X_i) = \theta, \text{Var}(X_i) = \theta(1-\theta)$$

$$f(x_i; \theta) = \theta^{x_i} \cdot (1-\theta)^{1-x_i}, \quad 0 \leq \theta \leq 1, x_i \in \{0, 1\}$$

Από το γεγονός  $x_i \in \{0, 1\} \rightarrow x_i = \theta + u_i$

$$(i) \prod_{i=1}^n f(x_i; \theta) = f(x_1; \theta) \cdot f(x_2; \theta) \dots f(x_n; \theta)$$

$$= \theta^{x_1} (1-\theta)^{1-x_1} \cdot \theta^{x_2} (1-\theta)^{1-x_2} \dots \theta^{x_n} (1-\theta)^{1-x_n}$$

$$= \theta^{\sum_{i=1}^n x_i} (1-\theta)^{\sum_{i=1}^n (1-x_i)} = \theta^{\sum_{i=1}^n x_i} (1-\theta)^{\sum_{i=1}^n 1 - \sum_{i=1}^n x_i}$$

$$= \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i} = L(\theta; x_i)$$

$$\ln L = \ln \left[ \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i} \right] =$$

$$= \ln \left[ \theta^{\sum_{i=1}^n x_i} \right] + \ln \left[ (1-\theta)^{n - \sum_{i=1}^n x_i} \right] =$$

$$= \sum_{i=1}^n x_i \cdot \ln \theta + \left( n - \sum_{i=1}^n x_i \right) \cdot \ln (1-\theta)$$

$$\underline{\text{F.O.C.}}: \frac{\partial \ln L}{\partial \theta} = \frac{1}{\theta} \sum_{i=1}^n x_i - \frac{1}{1-\theta} \left[ n - \sum_{i=1}^n x_i \right] = 0$$

$$\theta(1-\theta) \left( (1-\theta) \sum_{i=1}^n x_i - \theta \left[ n - \sum_{i=1}^n x_i \right] \right) = 0$$

$$(i) \quad \cancel{\sum_{i=1}^n x_i} - \theta \cancel{\sum_{i=1}^n x_i} - \theta n + \theta \cancel{\sum_{i=1}^n x_i} = 0 \quad \Leftrightarrow \quad \boxed{\hat{\theta} = \frac{1}{n} \sum_{i=1}^n x_i}$$

$$\text{S.O.C: } \frac{\partial^2 \ln L}{\partial \theta^2} = -\frac{1}{\theta^2} \sum_{i=1}^n x_i - \frac{1}{(1-\theta)^2} \left[ n - \sum_{i=1}^n x_i \right] \quad (A)$$

$$= -\frac{1}{\theta^2} n\theta - \frac{1}{(1-\theta)^2} (n - n\theta) =$$

$$= -\frac{n}{\theta} - \frac{1}{(1-\theta)^2} n(1-\theta) = -\frac{n}{\theta} - \frac{n}{1-\theta}$$

$$= \frac{-n(1-\theta) - n\theta}{\theta(1-\theta)} = \frac{-n + n\theta - n\theta}{\theta(1-\theta)} = \frac{-n}{\theta(1-\theta)} < 0$$

αρα  $\theta \in [0, 1]$

(ii) Απεκρίσιμη

$$\bullet E(\hat{\theta}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n E(x_i) =$$

$$= \frac{1}{n} \sum_{i=1}^n \theta = \frac{1}{n} n\theta = \theta$$

Άρα απεκρίσιμη

Συνέχεια

$$\text{Var}(\hat{\theta}) = \text{Var}\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(x_i)$$

$$= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(x_i) = \frac{1}{n^2} \sum_{i=1}^n \theta(1-\theta)$$

$$= \frac{1}{n^2} n\theta(1-\theta) = \frac{\theta(1-\theta)}{n}$$

## Γραφές Σωθικές

(5)

$$\lim_{n \rightarrow \infty} E(\hat{\theta}) = \lim_{n \rightarrow \infty} \theta = \theta \quad \checkmark$$

$$\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}) = \lim_{n \rightarrow \infty} \left[ \frac{\theta(1-\theta)}{n} \right] = 0 \quad \checkmark$$

Άρα  $n$  επαρκή για να είναι σωθικός.

## Πληροφορίες Διαφορετικές

$$C-R = I_n^{-1}(\theta) = \left[ E \left( - \frac{\partial^2 \ln L}{\partial \theta^2} \right) \right]^{-1} =$$

$$= \left[ E \left[ - \left( - \frac{n}{\theta(1-\theta)} \right) \right] \right]^{-1} =$$

$$= \left[ E \left( \frac{n}{\theta(1-\theta)} \right) \right]^{-1} = \frac{\theta(1-\theta)}{n} = \text{Var}(\hat{\theta})$$

Άρα ο εκτιμητής είναι πληροφ. διαφορετικός